

DRAFT REVISION OF IEEE STD 1139-1988 STANDARD DEFINITIONS OF PHYSICAL QUANTITIES FOR FUNDAMENTAL FREQUENCY AND TIME METROLOGY - RANDOM INSTABILITIES*

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Abstract - This is a draft revision of IEEE Std 1139-1988 Standard Definitions of Physical Quantities for Fundamental Frequency and Time Metrology. This draft standard covers the fundamental metrology for describing random instabilities of importance to frequency and time metrology. Quantities covered include frequency, amplitude, and phase instabilities; spectral densities of frequency, amplitude, and phase fluctuations; variances of frequency and phase fluctuations; time prediction; and confidence limits when estimating the variance from a finite data set. The standard unit of measure for characterizing phase and frequency instabilities in the frequency domain is $\mathcal{L}(f)$, defined as one half of the double-sideband spectral density of phase fluctuations. In the time domain, the standard unit of measure of frequency and phase instabilities is the fully overlapped Allan deviation $\sigma_y(\tau)$ or the fully overlapped modified Allan deviation $\text{Mod } \sigma_y(\tau)$.

1. Introduction

This is a draft revision of IEEE Std 1139-1988 Standard Definitions of Physical Quantities for Fundamental Frequency and Time Metrology which had been prepared by a previous SCC 27 consisting of Helmut Hellwig, Chairman; David Allan; Peter

Kartaschoff; Jacques Vanier; John Vig; Gernot M.R. Winkler; and Nicholas Yannoni. Some sections of the 1988 standard have remained unchanged. This standard covers the fundamental metrology for describing random instabilities of importance to frequency and time metrology. Quantities covered include frequency, amplitude, and phase instabilities; spectral densities of frequency, amplitude, and phase fluctuations; variances of frequency and phase fluctuations; time prediction; and confidence limits when estimating the variance from a finite data set. In addition, recommendations are made for the reporting of measurements of frequency and phase instabilities, especially as regards the recording of experimental parameters, experimental conditions and calculation techniques. This standard also covers translation between the frequency domain and time domain terminology and provides an extensive bibliography of the relevant literature. Systematic instabilities, such as environmental effects and aging, are discussed in IEEE Std 1193-1994.

2. Measures of Frequency, Amplitude and Phase Instabilities

The instantaneous output voltage of a precision oscillator can be expressed as

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$$V(t) = (V_o + \varepsilon(t)) \sin(2\pi\nu_o t + \phi(t)), \quad (\text{Eq. 1})$$

where V_o is the nominal peak voltage amplitude, $\varepsilon(t)$ is the deviation from the nominal amplitude, ν_o is the nominal frequency, and $\phi(t)$ is the phase deviation from the nominal phase $2\pi\nu_o t$. Figure 1 illustrates a signal with frequency, amplitude, and phase instabilities. As shown, frequency instability is the result of fluctuations in the period of oscillation. Fluctuations in the phase result in instability of the zero crossing. Fluctuations in the peak value of the signal (V_{peak}) result in amplitude instability.

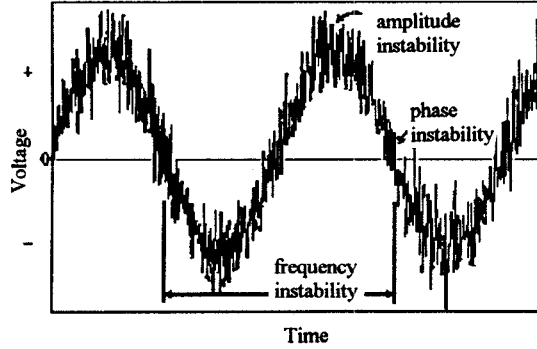


Figure 1. Instantaneous output voltage of an oscillator.¹

Frequency instability of a precision oscillator is defined in terms of the instantaneous, normalized frequency deviation, $y(t)$, as follows

$$y(t) = \frac{\nu(t) - \nu_o}{\nu_o} = \frac{\dot{\phi}(t)}{2\pi\nu_o}, \quad (\text{Eq. 2})$$

where $\nu(t)$ is the instantaneous frequency (time derivative of the phase divided by 2π), and

$$\dot{\phi}(t) = \frac{d\phi(t)}{dt}. \quad (\text{Eq. 3})$$

Amplitude instability is defined in terms of the instantaneous, normalized amplitude deviation

$$a(t) = \varepsilon(t)/V_o. \quad (\text{Eq. 4})$$

¹ In the signal shown the frequency components of the noise are higher than the carrier frequency. This is for illustration purposes only. In general, this standard applies to the frequency components of amplitude, phase and frequency instabilities which are lower in frequency than the carrier frequency.

Phase instability, defined in terms of the instantaneous phase deviation $\phi(t)$, can also be expressed in units of time, as

$$x(t) = \phi(t)/2\pi\nu_o. \quad (\text{Eq. 5})$$

With this definition, the instantaneous, normalized frequency deviation is

$$y(t) = dx(t)/dt. \quad (\text{Eq. 6})$$

Other random phenomena observed in certain oscillators are frequency jumps, that is, discontinuities in the frequency of oscillation. These phenomena are not repetitive or well understood and thus cannot be characterized by standard statistical methods.

3. Characterization of Frequency, Amplitude and Phase Instabilities

3.1 Frequency Domain. In the frequency domain, frequency, amplitude and phase instabilities can be defined or measured by one-sided spectral densities.

The unit of measure of frequency instability is the spectral density of fractional frequency fluctuations, $S_y(f)$, given by

$$S_y(f) = y^2(f) \frac{1}{BW}, \quad (\text{Eq. 7})$$

where $y(f)$ is the root mean squared fractional frequency deviation as a function of Fourier frequency, BW is the measurement system bandwidth in Hz, and the units of $S_y(f)$ are 1/Hz.

The unit of measure of amplitude instability is the spectral density of fractional amplitude fluctuations, $S_a(f)$, given by

$$S_a(f) = \left(\frac{\varepsilon(f)}{V_o} \right)^2 \frac{1}{BW}. \quad (\text{Eq. 8})$$

The units of $S_a(f)$ are 1/Hz.

Phase instability can be characterized by the spectral density of phase fluctuations, $S_\phi(f)$, given by

$$S_\phi(f) = \phi^2(f) \frac{1}{BW}. \quad (\text{Eq. 9})$$

The units of $S_\phi(f)$ are rad^2/Hz .

These spectral densities, $S_y(f)$, $S_a(f)$, and $S_\phi(f)$, are one-sided since the Fourier frequency f ranges from 0 to ∞ ; nevertheless, they include fluctuations from both the upper and the lower sidebands of the carrier.

$S_\phi(f)$ is the quantity that is generally measured in frequency metrology; however, $\mathcal{L}(f)$ (pronounced *script ell* of f) has become the prevailing measure of phase noise among manufacturers and users of frequency standards. According to the *old* definition [Kartaschoff, 1978], $\mathcal{L}(f)$ is the ratio of the power in one sideband due to phase modulation by noise (for a 1 Hz bandwidth) to the total signal power (carrier plus sidebands); that is,

$$\mathcal{L}(f) = \frac{\text{power density in one phase noise modulation sideband, per Hz}}{\text{total signal power}}. \quad (\text{Eq. 10})$$

Usually $\mathcal{L}(f)$ is expressed in decibels (dB) as 10 log ($\mathcal{L}(f)$), and its units are dB below the carrier in a 1 Hz bandwidth, generally abbreviated as dBc/Hz.

The *old* definition of $\mathcal{L}(f)$ is related to $S_\phi(f)$ by

$$\mathcal{L}(f) \cong \frac{S_\phi(f)}{2}. \quad (\text{Eq. 11})$$

This definition breaks down when the mean squared phase deviation, $\langle \phi^2(t) \rangle$ = the integral of $S_\phi(f)$ from f to ∞ , exceeds about 0.1 rad² or whenever there is a correlation between the power in the upper and lower sidebands. To circumvent this difficulty, $\mathcal{L}(f)$ is redefined as

$$\mathcal{L}(f) \equiv \frac{S_\phi(f)}{2}. \quad (\text{Eq. 12})$$

This redefinition is intended to avoid difficulties in the use of $\mathcal{L}(f)$ in situations where the small angle approximation is not valid and the correlation between the upper and lower sidebands is not measured. $\mathcal{L}(f)$, as defined by Eq. (12), is hereby designated as the standard measure of phase instability in the frequency domain. **The reasons are that (1) it can always be measured unambiguously, and (2) it conforms to the prevailing usage.**

Phase instability has sometimes been expressed in units of time by $S_x(f)$, the one-sided spectral density of the phase fluctuations expressed in units of time ($x(t)$),

$$S_x(f) = x^2(f) \frac{1}{\text{BW}}. \quad (\text{Eq. 13})$$

From Eq. (5), $S_\phi(f)$ and $S_x(f)$ are related by

$$S_x(f) = \frac{1}{(2\pi\nu_0)^2} S_\phi(f). \quad (\text{Eq. 14})$$

Since phase and frequency are directly related, that is, angular frequency is the time derivative of the phase, spectral densities of frequency and phase instabilities are also related:

$$S_y(f) = \frac{f^2}{\nu_0^2} S_\phi(f). \quad (\text{Eq. 15})$$

A device or signal should be characterized by a plot of $\mathcal{L}(f)$. In some applications, providing $\mathcal{L}(f)$ vs. discrete values of Fourier frequency is sufficient. (See Appendix A for further discussion.)

Other quantities related to phase instability are phase jitter and wander. Phase jitter is the integral of $S_\phi(f)$ over the Fourier frequencies of the application (usually above 10 Hz). When reporting phase jitter the range of Fourier frequencies and the slope of the discriminator used should be specified. Wander refers to the integral of $S_\phi(f)$ at Fourier frequencies below 10 Hz. It is possible to convert $S_\phi(f)$ to phase jitter by using the relation

$$\phi_{\text{jitter}}^2 = \int_{f_1}^{f_2} S_\phi(f) df, \quad (\text{Eq. 16})$$

where ϕ_{jitter}^2 refers to the phase jitter. It is not possible to obtain $S_\phi(f)$ from phase jitter, unless the shape of $S_\phi(f)$ is known.

3.2 Time-Domain. In the time domain, sequential average frequency instabilities are defined by a two-sample deviation $\sigma_y(\tau)$, also called the Allan deviation, which is the square root of a two-sample variance $\sigma_y^2(\tau)$, also called the Allan variance. This variance $\sigma_y^2(\tau)$ has no account of dead-time between adjacent frequency samples. (Dead time refers to the time between time-ordered data sets when no

measurement of frequency is taken.) For the sampling interval τ

$$\sigma_y(\tau) = \left[\frac{1}{2} \left\langle (\bar{y}(t+\tau) - \bar{y}(t))^2 \right\rangle \right]^{1/2} \quad (\text{Eq. 17})$$

$$= \left[\frac{1}{2} \left\langle (\bar{y}_{k+1} - \bar{y}_k)^2 \right\rangle \right]^{1/2},$$

where

$$\bar{y}_k = \frac{1}{\tau} \int_{t_k}^{t_k+\tau} y(t) dt = \frac{x(t_k+\tau) - x(t_k)}{\tau} = \frac{x_{k+1} - x_k}{\tau}. \quad (\text{Eq. 18})$$

The symbol $\langle \rangle$ denotes an infinite time average, and τ is the sampling interval. In practice, the requirement of infinite time average is never fulfilled, and the Allan deviation is estimated by

$$\sigma_y(\tau) \cong \left[\frac{1}{2(M-1)} \sum_{k=1}^{M-1} (\bar{y}_{k+1} - \bar{y}_k)^2 \right]^{1/2}, \quad (\text{Eq. 19})$$

where M is the number of frequency measurements.

The Allan deviation can also be expressed in terms of time difference (or time residual) measurements by combining Eqs. 18 and 19:

$$\sigma_y(\tau) \cong \left[\frac{1}{2(N-2)\tau^2} \sum_{k=1}^{N-2} (x_{k+2} - 2x_{k+1} + x_k)^2 \right]^{1/2}, \quad (\text{Eq. 20})$$

where x_k , x_{k+1} , and x_{k+2} are time residual measurements made at t_k , $t_{k+1} = t_k + \tau$, and $t_{k+2} = t_k + 2\tau$, $k=1, 2, 3, \dots$, N is the number of time measurements, and $1/\tau$ is the nominal fixed sampling rate which gives zero dead time between frequency measurements. *Residual* implies the consistent, systematic effects such as frequency drift have been removed (see section 3.3).

If there is dead time between the frequency departure measurements and it is ignored in the computation of $\sigma_y(\tau)$, resulting instability values will be biased (except for white frequency noise). Some of the biases have been studied and some correction tables published [Barnes, 1969; Lesage, 1983; Barnes and Allan, 1988]. Therefore, the term $\sigma_y(\tau)$

shall not be used to describe such biased measurements without stating the bias together with $\sigma_y(\tau)$. The unbiased $\sigma_y(\tau)$ can be calculated from the biased, using information in the references. Considering that $\{x_k\}$ can be routinely measured, it is preferred that $\{x_k\}$ is used to compute $\sigma_y(\tau)$ since the problem of dead-time is solved.

If the initial sampling rate is specified as $1/\tau_0$, then, in general, we may obtain an estimate of $\sigma_y(\tau)$ with better confidence using what is called *overlapping estimates*. This estimate is obtained by computing

$$\sigma_y(\tau) = \left[\frac{1}{2(N-2m)\tau^2} \sum_{k=1}^{N-2m} (x_{k+2m} - 2x_{k+m} + x_k)^2 \right]^{1/2}, \quad (\text{Eq. 21})$$

where N is the number of original time residual measurements spaced by τ_0 ($N = M + 1$, where M is the number of original frequency measurements of sample time τ_0) and $\tau = m\tau_0$. Examples of overlapped $\sigma_y(\tau)$ estimates are given in Appendix B.

Equation 21 shows that $\sigma_y(\tau)$ acts like a second-difference operator on the time deviation residuals usually providing a stationary measure of the stochastic behavior even for nonstationary processes. An efficient spacing of τ values in a plot of $\log[\sigma_y(\tau)]$ vs. $\log[\tau]$ sets $m = 2^p$, where $p = 0, 1, 2, 3, \dots$.

The frequency stability of an oscillator should be characterized by a plot of $\sigma_y(\tau)$ vs. sampling time (τ). In some applications, providing discrete values of $\sigma_y(\tau)$ vs. sampling time τ is sufficient. The measurement system bandwidth should always be specified. (See Appendix A for further discussion.)

When differentiating between white and flicker PM noise is desirable, a *modified deviation*, denoted as $\text{Mod } \sigma_y(\tau)$, may be used to characterize frequency instabilities [Allan and Barnes, 1981; Stein, 1985]. Unlike $\sigma_y(\tau)$, $\text{Mod } \sigma_y(\tau)$ has the property of yielding different dependence on τ for white phase noise and flicker phase noise; the dependencies are $\tau^{-3/2}$ and τ^{-1} , respectively. (The dependence for $\sigma_y(\tau)$ is τ^{-1} for both white and flicker phase noise.) Another advantage is that $\text{Mod } \sigma_y(\tau)$ averages wideband phase noise faster than τ^{-1} . $\text{Mod } \sigma_y(\tau)$ is defined as

$$\text{Mod } \sigma_y(\tau) = \left\{ \frac{1}{2\tau^2 m^2 (N-3m+1)} \sum_{j=1}^{N-3m+1} \left[\sum_{i=j}^{m+j-1} (x_{i+2m} - 2x_{i+m} + x_i) \right]^2 \right\}^{1/2}. \quad (\text{Eq. 22})$$

For examples of $\sigma_y(\tau)$ and $\text{Mod } \sigma_y(\tau)$ see Appendix B.

A unit of measure that is often used in time transfer systems, such as GPS, is $\sigma_x(\tau)$. $\sigma_x(\tau)$ is defined as

$$\sigma_x(\tau) = \frac{\tau}{\sqrt{3}} \text{Mod } \sigma_y(\tau). \quad (\text{Eq. 23})$$

This quantity is useful when white and flicker of phase modulation noise dominate a synchronization system.

Another approach to distinguish different noise types is to use multivariate analysis [Vernotte, et al., 1993]. By using several variances to analyze the same data it is possible to estimate the coefficients for the 5 noise types. For a description of noise types see Appendix A.

At long averaging times, when the number of samples is small ($N \leq 5$), $\sigma_y(\tau)$ has a bias related to its insensitivity to noise processes which appear as “odd” or antisymmetric functions over a finite measurement of $x(t)$ (odd with respect to the middle point of the data). An extension of the Allan deviation ($\hat{\sigma}_{y,\text{TOTAL}}(\tau)$) which circularizes the original data provides a better estimate of frequency stability which is most noticeable at long averaging times and is recommended [Howe, 1995]. Details on $\hat{\sigma}_{y,\text{TOTAL}}(\tau)$ are given in Appendix B.

3.3 Systematic Instabilities. The long-term frequency change of a source is called frequency drift. Drift includes frequency changes caused by changes in the components of the oscillator, in addition to sensitivities to the oscillator’s changing environment and changes caused by load and power supply changes [Vig and Meeker, 1991].

The frequency aging of an oscillator refers to the change in the frequency of oscillation caused by changes in the components of the oscillator, either in the resonant unit or in the accompanying electronics. Aging differs from drift in that it does

not include frequency changes due to changes in the environment such as temperature changes. Aging is thus a measure of the long-term stability of the oscillator, independent of its environment. The frequency aging of a source (positive or negative) is typically maximum immediately after turn-on.

Aging can be specified by the normalized rate of change in frequency at a specified time after turn-on (for example, 1×10^{-10} per day after 30 days), or by the total normalized change in frequency in a period of time (for example, 1×10^{-8} per month) [Vig and Meeker, 1991].

3.4 Clock-Time Prediction. The variation of the time difference between a real clock and an ideal uniform time scale, also known as time interval error TIE, observed over a time interval starting at time t_0 and ending at $t_0 + t$ is defined as

$$\text{TIE}(t) = x(t_0 + t) - x(t_0) = \int_{t_0}^{t_0+t} y(t') dt'. \quad (\text{Eq. 24})$$

For fairly simple models, regression analysis can provide efficient estimates of the TIE [Draper and Smith, 1966; CCIR, 1986]. In general, there are many estimators possible for any statistical quantity. Ideally, we would like an efficient and unbiased estimator. Using the time domain measure $\sigma_y(\tau)$ defined in section 3.2, the following estimate of the standard deviation (RMS) of TIE and its associated systematic departure due to a linear frequency drift (plus its uncertainty) can be used to predict a probable time interval error of a clock synchronized at $t = t_0 = 0$ and left free running thereafter

$$\text{RMS TIE}_{\text{est}}(t) = t \left(\left[\frac{x(t_0)}{t} \right]^2 + \sigma_{y_0}^2 + \sigma_y^2(\tau = t) + \frac{a^2}{4} t^2 \right)^{1/2}, \quad (\text{Eq. 25})$$

where

- $x(t_0)$ = initial synchronization uncertainty,
- σ_{y_0} = two-sample deviation of the initial frequency adjustment,
- $\sigma_y(\tau)$ = the two-sample deviation describing the random frequency instability of the clock at $\tau = t$,

a = normalized linear frequency drift per unit of time plus the uncertainty in the drift estimate.

The third term in the brackets provides an optimum and unbiased estimate (under the condition of an optimum (RMS) prediction method) in the cases of white noise FM and/or random walk FM. The third term is too optimistic, by about a factor of 1.4, for flicker noise FM, and too pessimistic, by about a factor of 3, for white noise PM.

This estimate is a useful and fairly simple approximation. A more complete error analysis becomes difficult; if carried out, such an analysis needs to include the methods of time prediction, the uncertainties of the clock parameters using the confidence limits of measurements defined below, the detailed clock noise models, systematic effects, etc.

A quantity often used to characterize the stability of clocks in telecommunication systems is the maximum time interval error (MTIE). MTIE is defined as the maximum time difference between a clock and an ideal reference [Bregni, 1996].

4. Confidence Limits of Measurements

A simple method to compute the confidence interval for $\sigma_y(\tau)$ [Lesage and Audoin, 1973], which assumes a symmetric (Gaussian) distribution, uses the relation

$$I_\alpha \cong \sigma_y(\tau) \kappa_\alpha M^{-1/2}, \quad (\text{Eq. 26})$$

where I_α is the confidence interval, κ is a constant, α is an integer that depends on the type of noise (see Appendix A), and M is the total number of data points used in the estimate. For a 1σ or 68 % confidence interval the values for κ_α are

$$\begin{aligned} \kappa_2 &= 0.99, \\ \kappa_1 &= 0.99, \\ \kappa_0 &= 0.87, \\ \kappa_{-1} &= 0.77, \\ \kappa_{-2} &= 0.75. \end{aligned}$$

As an example of the Gaussian model with $M = 100$, $\alpha = -1$ (flicker frequency noise) and $\sigma_y(\tau = 1 \text{ s}) = 1 \times 10^{-12}$, we may write

$$I_\alpha \cong \sigma_y(\tau)(0.77)(100)^{-1/2} = \sigma_y(\tau) (0.077),$$

which gives

$$\sigma_y(\tau = 1 \text{ s}) = (1 \pm 0.08) 10^{-12}.$$

This analysis for $\sigma_y(\tau)$ is valid only for $M \geq 10$. If M is small, then the plus and minus confidence intervals become sufficiently asymmetric and the κ_α coefficients are not valid; however, these confidence intervals can be calculated [Lesage and Audoin, 1973].

Another way of computing confidence intervals for $\sigma_y(\tau)$ is to use the chi-squared distribution. The estimated Allan variance has a chi-squared distribution function given by Eq. 27. The number of degrees of freedom for a specific noise process and number of samples can be computed and then used in Eq. 27 to compute the confidence interval [Howe, 1981]:

$$\chi^2 = (\text{df}) \frac{\hat{\sigma}_y^2}{\sigma_y^2}, \quad (\text{Eq. 27})$$

where df is the number of degrees of freedom, $\hat{\sigma}_y^2$ is the estimated (measured) Allan variance, and σ_y^2 is the true Allan variance. Table C1 (Appendix C) shows empirical equations to compute the number of degrees of freedom for different types of noise processes (N number of samples, and $m = \tau/\tau_0$). Examples of this method are shown in Appendix C.

Other methods have been developed for calculating confidence intervals for $\text{Mod } \sigma_y(\tau)$. For a detailed description see [Walter, 1994; Greenhall, 1995]. Comparison of confidence intervals for $\sigma_y(\tau)$ (no-overlap, full-overlap) and $\text{Mod } \sigma_y(\tau)$ are given in Appendix C (Table C2)].

5. Standards for Characterizing or Reporting Measurements of Frequency and Phase Instabilities

The standard unit of measure for characterizing phase and frequency instabilities in the frequency domain is $\mathcal{L}(f)$, defined as one half of the double-sideband spectral density of phase fluctuations (see Eq. 12). When expressed in decibels its units are dBc/Hz (dB below the carrier in a 1 Hz bandwidth). A device should be characterized by a plot of $\mathcal{L}(f)$ versus f . In the time domain, the standard unit of

measure of frequency and phase instabilities is the fully overlapped Allan deviation $\sigma_y(\tau)$ or the fully overlapped modified Allan deviation $\text{Mod } \sigma_y(\tau)$. A device should be characterized by a plot of $\sigma_y(\tau)$ or $\text{Mod } \sigma_y(\tau)$ versus τ . The measurement system bandwidth (f_h) and the total measurement time should be indicated.

In addition the following provisions are recommended when reporting measurements on frequency and phase instabilities.

5.1 Nonrandom phenomena should be recognized; for example:

- (1) Any observed time dependence of the statistical measures should be stated.
- (2) The method of modeling systematic behavior should be specified (for example, an estimate of the linear frequency drift was obtained from the coefficients of a linear least-squares regression to M frequency measurements, each with a specified averaging or sample time t and measurement bandwidth f_h).
- (3) The environmental sensitivities should be stated (for example, the dependence of frequency and/or phase on temperature, magnetic field, barometric pressure, vibration, etc.).

5.2 Relevant measurement or specification parameters should be given:

- (4) The method of measurements.
- (5) The characteristics of the reference signal (equal noise or much lower noise assumed).
- (6) The nominal signal frequency ν_0 .
- (7) The measurement system bandwidth f_h and the corresponding low pass filter response.
- (8) The total measurement time and the number of measurements M .
- (9) The calculation techniques (for example, details of the window function when estimating power spectral densities from time domain data, or the assumptions about effects of dead time when estimating the two-sample deviation $\sigma_y(t)$).
- (10) The confidence of the estimate (or error bar) and its statistical probability (for example, 1σ for 68%, 2σ for 95%).
- (11) The environment during measurement.
- (12) If a passive element, such as a crystal filter, is being measured in contrast to a frequency and/or time generator.

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7. References

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Appendix A

Power-Laws and Conversion between Frequency and Time Domain

A1. Power-Law Spectral Densities. Power-law spectral densities serve as reasonable and accurate models of the random fluctuations in precision oscillators. In practice, these random fluctuations can often be represented by the sum of five such noise processes assumed to be independent, as

$$S_y(f) = \begin{cases} \sum_{\alpha=-2}^{+2} h_{\alpha} f^{\alpha} & \text{for } 0 < f < f_h, \\ 0 & \text{for } f \geq f_h, \end{cases} \quad (\text{Eq. A1})$$

where

h_{α} = constant,

α = integer,

f_h = high-frequency cut-off of an infinitely sharp low pass filter.

High frequency divergence is eliminated by the restrictions on f in this equation. The identification and characterization of the five noise processes are given in Table A1, and shown in Fig. A1.

A2. Conversion between Frequency and Time Domain. The operation of the counter, averaging the frequency for a time τ , may be thought of as a filtering operation. The transfer function $H(f)$ of this equivalent filter is then the Fourier transform of the impulse response of the filter. The time domain frequency instability is then given by

$$\sigma^2(M, T, \tau) = \int_0^{\infty} S_y(f) |H(f)|^2 df, \quad (\text{Eq. A2})$$

where $S_y(f)$ is the spectral density of frequency fluctuations. $1/T$ is the measurement rate ($T-\tau$ is the dead time between measurements). In the case of the two-sample variance $|H(f)|^2$ is $2(\sin^4 \pi \tau f) / (\pi \tau f)^2$. The two-sample variance can thus be computed from

$$\sigma_y^2(\tau) = 2 \int_0^{f_h} S_y(f) \frac{\sin^4(\pi \tau f)}{(\pi \tau f)^2} df. \quad (\text{Eq. A3})$$

Specifically, for the power law model given, the time domain measure also follows a power law:

$$\sigma_y^2(\tau) = h_{-2} \frac{(2\pi)^2}{6} \tau + h_{-1} 2 \log_e 2 + h_0 \frac{1}{2\tau} + h_1 \frac{1.038 + 3 \log_e(2\pi f_h \tau)}{(2\pi)^2 \tau^2} + h_2 \frac{3f_h}{(2\pi)^2 \tau^2}. \quad (\text{Eq. A4})$$

Equation A4 assumes that f_h is the high frequency cut-off of an infinitely sharp low pass filter and that $2\pi f_h \tau \gg 1$. This equation also implicitly assumes that the random driving mechanism for each term is independent of the others, and that the mechanism is valid over all Fourier frequencies, which may not always be true.

The values of h_α are characteristic models of oscillator frequency noise. For integer values (as often seems to be the case for reasonable models),

$$\begin{aligned} \mu &= -\alpha - 1, \text{ for } -3 \leq \alpha \leq 1, \\ \mu &\approx -2 \text{ for } \alpha \geq 1, \end{aligned}$$

where

$$\sigma_y^2(\tau) \sim \tau^\mu.$$

The modified two-sample variance can also be computed from $S_y(f)$ by using

$$\text{Mod} \sigma_y^2(\tau) = \frac{2}{n^4} \int_0^{f_h} S_y(f) \frac{\sin^6(\pi \tau f)}{(\pi \tau_0 f)^2 \sin^2(\pi \tau_0 f)} df. \quad (\text{Eq. A5})$$

Table A2 gives the coefficients of the translation from $S_y(f)$ (frequency domain) to $\sigma_y^2(\tau)$ (time domain). In general computation of $S_y(f)$ or related frequency domain measurements from $\sigma_y(\tau)$ or $\text{Mod} \sigma_y(\tau)$ are not permitted unless only one power law noise type is present. Nevertheless, when several noise types are present, special analysis can be made on the time domain data to obtain the coefficients (in the frequency domain) for each power law [Vernotte, et al., 1993].

The slope characteristics of the five independent noise processes are plotted in the frequency and time domains in Fig. A1 (log-log scale).

Table A1
The Functional Characteristics of the Independent Noise Processes Used in Modeling Frequency Instability of Oscillators

Description of Noise Process	Slope Characteristics of Log Log Plot				
	Frequency Domain		Time-Domain		
	$S_y(f)$ or $S_{\dot{\phi}}(f)$	$S_{\phi}(f)$ or $S_x(f)$	$\sigma_y^2(\tau)$	$\sigma_y(\tau)$	Mod $\sigma_y(\tau)$
	α	β	μ	$\mu/2$	μ'
Random walk frequency modulation	-2	-4	1	1/2	1/2
Flicker frequency modulation	-1	-3	0	0	0
White frequency modulation	0	-2	-1	-1/2	-1/2
Flicker phase modulation	1	-1	-2	-1	-1
White phase modulation	2	0	-2	-1	-3/2

$$S_y(f) = \frac{(2\pi f)^2}{(2\pi\nu_0)^2} S_{\dot{\phi}}(f) = h_\alpha f^\alpha$$

$$S_{\dot{\phi}}(f) = \nu_0^2 h_\alpha f^{\alpha-2} = \nu_0^2 h_\alpha f^\beta \quad (\beta \equiv \alpha - 2)$$

$$S_x(f) = \frac{1}{4\pi^2} h_\alpha f^{\alpha-2} = \frac{1}{4\pi^2} h_\alpha f^\beta$$

$$\sigma_y^2(\tau) \sim |\tau|^\mu$$

$$\sigma_y(\tau) \sim |\tau|^{\mu/2}$$

$$\text{Mod} \sigma_y(\tau) \sim |\tau|^{\mu'}$$

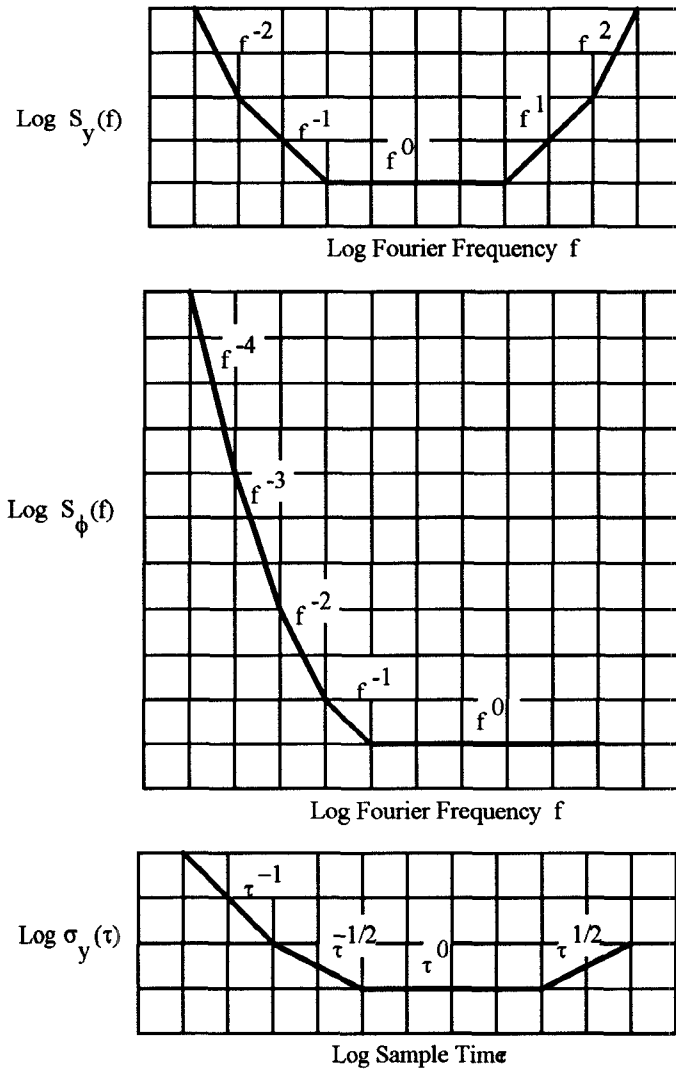


Fig. A1 Slope Characteristics of the Five Independent Noise Process

Table A2. Translation of Frequency Instability Measures from Spectral Densities in Frequency Domain to Variances in Time Domain for an Infinitely Sharp Low Pass Filter with $2\pi f_h \tau \gg 1$.

Description of Noise Process	$S_y(f) =$	$S_\phi(f) =$	$\sigma_y^2(\tau) =$
Random walk frequency modulation	$h_{-2}f^{-2}$	$h_{-2}v^2f^{-4}$	$Ah_{-2}\tau^1$
Flicker frequency modulation	$h_{-1}f^{-1}$	$h_{-1}v^2f^{-3}$	$Bh_{-1}\tau^0$
White frequency modulation	h_0f^0	$h_0v^2f^{-2}$	$Ch_0\tau^{-1}$
Flicker phase modulation	h_1f^1	$h_1v^2f^{-1}$	$Dh_1\tau^{-2}$
White phase modulation	h_2f^2	$h_2v^2f^0$	$Eh_2\tau^{-2}$

$$A = \frac{2\pi^2}{3}$$

$$B = 2 \ln 2$$

$$C = 1/2$$

$$D = \frac{1.038 + 3 \ln(2\pi f_h \tau)}{4\pi^2}$$

$$E = \frac{3f_h}{4\pi^2}$$

Appendix B

Examples and Additional Variances that may be used to Describe Frequency Instabilities in the Time Domain

This appendix contains basic examples on how to compute the variances used to describe frequency instabilities in the time domain. For more information on this topic and on how to assess the validity of the computations when using larger number of samples see [Riley, 1995; Riley, 1996].

B1. $\sigma_y(\tau)$ - Examples. Fig. B1 shows a plot of the time deviation between a pair of oscillators as a function of time. The recorded time samples for $\tau = 1$ s are shown in the first column of Table B1. To compute $\sigma_y(\tau = 1$ s) we need to compute the average fractional frequency deviation for x_k s separated by 1 s, then calculate the difference between adjacent \bar{y}_k s and use Eq. 19 to obtain $\sigma_y(\tau)$. See Table B1.

Table B1. Steps to compute $\sigma_y(1$ s).		
x_k (10^{-6})	$\bar{y}_k = (x_{k+1} - x_k)/\tau$ (10^{-6})	$\bar{y}_{k+1} - \bar{y}_k$ (10^{-6})
0	--	--
43.6	$\bar{y}_1 = (x_2 - x_1)/\tau = 43.6$	--
89.7	$\bar{y}_2 = (x_3 - x_2)/\tau = 46.1$	2.5
121.6	$\bar{y}_3 = (x_4 - x_3)/\tau = 31.9$	-14.2
163.7	$\bar{y}_4 = (x_5 - x_4)/\tau = 42.1$	10.2
208.4	$\bar{y}_5 = (x_6 - x_5)/\tau = 44.7$	2.6
248	$\bar{y}_6 = (x_7 - x_6)/\tau = 39.6$	-5.1
289	$\bar{y}_7 = (x_8 - x_7)/\tau = 41.0$	1.4
319.8	$\bar{y}_8 = (x_9 - x_8)/\tau = 30.8$	-10.2

In this example $N = 9$ (number of time samples) and $M = 8$; therefore

$$\sigma_y(\tau = 1\text{ s}) = \left[\frac{1}{2(7)} \sum_{k=1}^7 (\bar{y}_{k+1} - \bar{y}_k)^2 \right]^{1/2}, \quad (\text{Eq. B1})$$

$$\sigma_y(1\text{ s}) = [3.2 \times 10^{-11}]^{1/2} = 5.67 \times 10^{-6}.$$

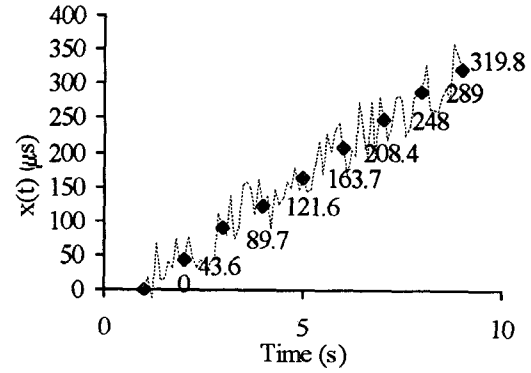


Figure B1. Plot of $x(t)$ between a pair of oscillators.

For $\tau = 2$ s ($= 2\tau_0$) the procedure is similar: compute the fractional frequency deviation for x_k s separated by 2 s, then calculate the difference between adjacent \bar{y}_k s and use Eq. 19 to obtain $\sigma_y(\tau)$. See Table B2.

Table B2. Steps for computing $\sigma_y(2$ s).		
x_k (10^{-6})	$\bar{y}_k = (x_{k+2} - x_k)/\tau$ (10^{-6})	$\bar{y}_{k+2} - \bar{y}_k$ (10^{-6})
0	--	--
43.6	--	--
89.7	$\bar{y}_1 = (x_3 - x_1)/2 = 44.85$	--
121.6	--	--
163.7	$\bar{y}_3 = (x_5 - x_3)/2 = 37$	-7.85
208.4	--	--
248	$\bar{y}_5 = (x_7 - x_5)/2 = 42.15$	5.15
289	--	--
319.8	$\bar{y}_7 = (x_9 - x_7)/2 = 35.9$	-6.25

For this example $M = 4$ since there are a total of four \bar{y}_k s. Therefore

$$\sigma_y(\tau = 2\text{ s}) = \left[\frac{1}{2(3)} \sum_{k=1}^3 (\bar{y}_{k+2} - \bar{y}_k)^2 \right]^{1/2}, \quad (\text{Eq. B2})$$

$$\sigma_y(2\text{ s}) = [2.12 \times 10^{-11}]^{1/2} = 4.6 \times 10^{-6}.$$

As mentioned in Section 3.2, it is usually more efficient to use overlapped estimates when possible since it results in a better confidence interval. Figure B2 illustrates how to compute the \bar{y}_k s for an overlapped estimate of $\sigma_y(\tau = 2$ s). In this case $m = 2$ ($\tau = 2\tau_0$) and $M = 8$. The \bar{y}_k s and the second difference values are shown in Table B3.

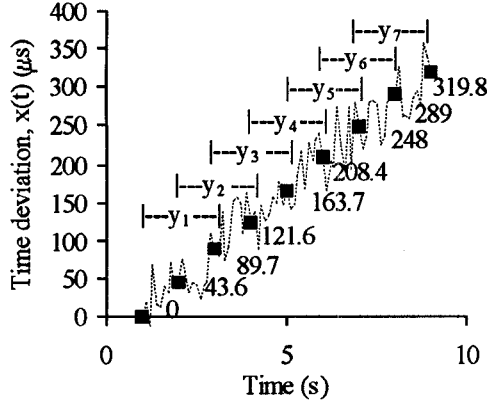


Figure B2. Computation of \bar{y}_k s for overlapped estimates.

There are a total of $(N - 2m = 5)$ second difference values $(\bar{y}_{k+m} - \bar{y}_k)$; therefore the Allan deviation equation becomes

$$\sigma_y(\tau) = \left[\frac{1}{2(N-2m)} \sum_{k=1}^{N-2m} (\bar{y}_{k+m} - \bar{y}_k)^2 \right]^{\frac{1}{2}}, \quad (\text{Eq. B3})$$

where $\bar{y}_k = (x_{k+m} - x_k)/\tau$. Equation B3 becomes Eq. 21 when the \bar{y}_k s are expressed in terms of the initial time residual measurements. It is used in this example to help explain the origin of Eq. 21. Using the values in Table B3 (last column), Eq. B3 yields

$$\sigma_y(2s) = \left\{ \frac{1}{2(5)} [(-7.85)^2 + 4.4^2 + 5.15^2 + (-3.1)^2 + (-6.25)^2] \right\}^{\frac{1}{2}},$$

$$\sigma_y(2s) = [1.56 \times 10^{-11}]^{\frac{1}{2}} = 3.95 \times 10^{-6}.$$

B2. Mod $\sigma_y(\tau)$ - Example. The Modified Allan deviation can also be used to characterize frequency stability in the time domain. This deviation uses the average of m adjacent x_k s when computing the stability for $\tau = m\tau_0$. The fractional frequency deviations are then obtained using the \bar{x}_k s. See Fig. B3 for computation of \bar{x}_k s and \bar{y}_k 's.

Table B4 shows the computed \bar{x}_k s, and \bar{y}_k 's for $\tau = 2s$. The Modified Allan deviation can then be obtained by using Eq. 19 and the fact that $m = 2$ and the equivalent M is $N - 3m + 1$:

$$\text{Mod } \sigma_y(\tau) = \left\{ \frac{1}{2(N-3m+1)} \sum_{k=1}^{N-3m+1} (\bar{y}'_{k+m} - \bar{y}'_k)^2 \right\}^{\frac{1}{2}}, \quad (\text{Eq. B4})$$

$$\text{Mod } \sigma_y(2s) = \left\{ \frac{1}{2(4)} [(-1.73)^2 + 4.78^2 + 1.03^2 + (-4.68)^2] \right\}^{\frac{1}{2}} = [6.1 \times 10^{-12}]^{\frac{1}{2}},$$

$$\text{Mod } \sigma_y(2s) = 2.47 \times 10^{-6}.$$

Note that Eq. B4 becomes Eq. 22 when expressing the \bar{y}_k 's in terms of the initial time residual measurements. It is used in this example to help explain the origin of Eq. 22.

Table B3. Steps for computing an overlapped estimate of $\sigma_y(2s)$.

x_k (10^{-6})	$\bar{y}_k = (x_{k+2} - x_k)/\tau$ (10^{-6})	$\bar{y}_{k+2} - \bar{y}_k$ (10^{-6})
0	--	--
43.6	--	--
89.7	$\bar{y}_1 = (x_3 - x_1)/\tau = 44.9$	--
121.6	$\bar{y}_2 = (x_4 - x_2)/\tau = 39$	--
163.7	$\bar{y}_3 = (x_5 - x_3)/\tau = 37$	-7.85
208.4	$\bar{y}_4 = (x_6 - x_4)/\tau = 43.4$	4.4
248	$\bar{y}_5 = (x_7 - x_5)/\tau = 42.2$	5.15
289	$\bar{y}_6 = (x_8 - x_6)/\tau = 40.3$	-3.1
319.8	$\bar{y}_7 = (x_9 - x_7)/\tau = 35.9$	-6.25

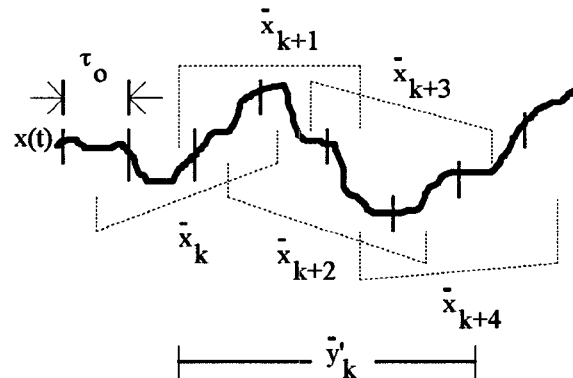


Figure B3. Method for calculating \bar{x}_k s, and \bar{y}_k 's for Mod $\sigma_y(\tau)$.

Table B4. Computed \bar{x}_k and \bar{y}_k' values for Mod $\sigma_y(2\text{ s})$.

x_k (10^{-6})	$\bar{x}_k = (x_{k+1} + x_k)/2$ (10^{-6})	$\bar{y}_k' = (\bar{x}_{k+2} - \bar{x}_k)/\tau$ (10^{-6})	$\bar{y}_{k+2}' - \bar{y}_k'$ (10^{-6})
0	--	--	--
43.6	21.8	--	--
89.7	66.65	--	--
121.6	105.65	$\bar{y}_1' = (\bar{x}_3 - \bar{x}_1)/2 = 41.93$	--
163.7	142.65	$\bar{y}_2' = (\bar{x}_4 - \bar{x}_2)/2 = 38$	--
208.4	186.05	$\bar{y}_3' = (\bar{x}_5 - \bar{x}_3)/2 = 40.2$	-1.73
248	228.2	$\bar{y}_4' = (\bar{x}_6 - \bar{x}_4)/2 = 42.78$	4.78
289	268.5	$\bar{y}_5' = (\bar{x}_7 - \bar{x}_5)/2 = 41.23$	1.03
319.8	304.4	$\bar{y}_6' = (\bar{x}_8 - \bar{x}_6)/2 = 38.1$	-4.68

B3. Extension of the Allan Deviation $\hat{\sigma}_{y,\text{TOTAL}}(\tau)$. In section 3.2 we noted that when the number of samples is small ($N \leq 5$), the Allan deviation has a bias related to its insensitivity to odd (antisymmetric) noise processes in $x(t)$ (odd about the midpoint of the $x(t)$ data or even in terms of average \bar{y}_k). This is illustrated in Fig. B4. Figure B4.1 shows three phase samples of a noise process which is odd about x_2 . The calculated fractional frequency deviations according to Eq. 19 are shown in Fig. B4.2. Since \bar{y}_1 and \bar{y}_2 are equal, contributions due to this noise process will not show up in $\sigma_y(\tau)$.

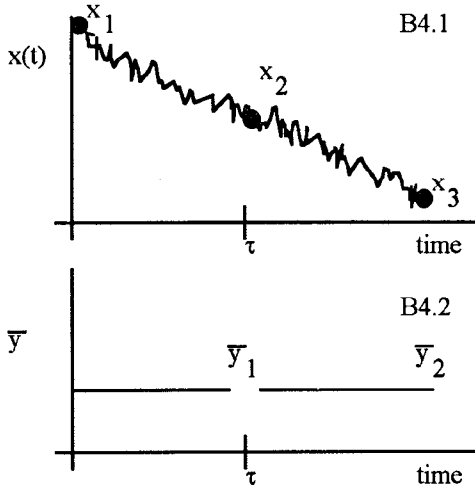


Figure B4. Odd noise process about x_2 .

For this reason, when $N \leq 5$, we recommend using an extension of the 2-sample (Allan) deviation which adds the odd (3-sample) component [Howe, 1995]. The estimate to this deviation, $\hat{\sigma}_{y,\text{TOTAL}}(\tau)$, wraps the x_k data to provide a better estimate of frequency stability. The advantages of $\hat{\sigma}_{y,\text{TOTAL}}(\tau)$ are outlined in several references [Howe, 1995; Howe, 1997] and its generalized form is given by

$$\hat{\sigma}_{y,\text{TOTAL}}(\tau) = \left(\frac{1}{\tau^2} \frac{1}{2(N-m-1)} \right)^{1/2} \left(\sum_{k=1}^{N-m-1} [x'_{k+2m} - 2x'_{k+m} + x'_k]^2 \right)^{1/2}, \quad (\text{Eq. B5})$$

where the x'_k s represent the time residual measurements after the linear trend (or slope in phase) has been removed (matching the endpoints of the data set $\{x_k\}$). $\hat{\sigma}_{y,\text{TOTAL}}(\tau)$ can also be represented in terms of fractional frequency fluctuation averages as

$$\hat{\sigma}_{y,\text{TOTAL}}(\tau) = \left\{ \frac{1}{N-2m+1} \sum_j^{N-2m+1} \left[\frac{1}{2(N-2m)} \sum_{k=1}^{N-2m} (\bar{y}_{k+m,j} - \bar{y}_{k,j})^2 \right] \right\}^{1/2}, \quad (\text{Eq. B6})$$

where the series of $\bar{y}_{k,j}$ (fractional frequency fluctuations averages) goes as $\bar{y}_{j+1}, \bar{y}_{j+2}, \dots, \bar{y}_{N-1}, \bar{y}_1, \dots, \bar{y}_j$.

As an example we will compute $\hat{\sigma}_{y,TOTAL}(2s)$ for $x(t)$ in Fig. B5. Notice that this data set is different from the one used in the previous examples. In this case x_1 , x_3 , and x_5 almost fall into a line; therefore the value for $\sigma_y(\tau)$ will be negatively biased (too optimistic).

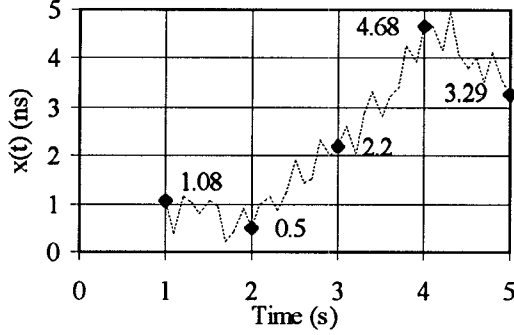


Figure B5. $x(t)$ as a function of time.

Starting with the time difference measurements $\{x_k\}$, the first step is to compute the overall frequency difference (OFD = $(x_N - x_1)/(N-1)\tau_0$) and remove it from each x_k , the result being that x_1 and x_5 will have the same value. This is necessary in order to be able to wrap the data without a discontinuous phase step ($x_5 \rightarrow x_1$). Therefore,

$$OFD = \frac{x_5 - x_1}{4} = \frac{3.29 - 1.08}{4} = 0.5525 \times 10^{-9}. \quad (\text{Eq. B7})$$

Now remove the OFD from x_k s:

$$\begin{aligned} x'_1 &= 1.08 \text{ ns}, \\ x'_2 &= 0.5 - OFD = -0.0525 \text{ ns}, \\ x'_3 &= 2.2 - 2(OFD) = 1.095 \text{ ns}, \\ x'_4 &= 4.68 - 3(OFD) = 3.0225 \text{ ns}, \\ x'_5 &= 3.29 - 4(OFD) = 1.08 \text{ ns}. \end{aligned}$$

$\hat{\sigma}_{y,TOTAL}(2s)$ can then be computed using Eq. B8:

$$\hat{\sigma}_{y,TOTAL}(2 \text{ sec}) = \left\{ \frac{1}{2(2)} [(\bar{y}_{2,1} - \bar{y}_{1,1})^2 + (\bar{y}_{2,2} - \bar{y}_{1,2})^2] \right\}^{\frac{1}{2}}, \quad (\text{Eq. B8})$$

where

$$\begin{aligned} \bar{y}_{1,1} &= \frac{x'_3 - x'_1}{2} = \frac{1.095 - 1.08}{2} = 0.0075 \times 10^{-9}, \\ \bar{y}_{2,1} &= \frac{x'_5 - x'_3}{2} = \frac{1.08 - 1.095}{2} = -0.0075 \times 10^{-9}, \\ \bar{y}_{1,2} &= \frac{x'_4 - x'_2}{2} = \frac{3.0225 + 0.0525}{2} = 1.5375 \times 10^{-9}, \\ \bar{y}_{2,2} &= \frac{x'_2 - x'_4}{2} = \frac{-0.0525 - 3.0225}{2} = -1.5375 \times 10^{-9}. \end{aligned}$$

Therefore,

$$\begin{aligned} \hat{\sigma}_{y,TOTAL}(2s) &= \left\{ \frac{1}{4} [(-0.015)^2 + (-3.075)^2] \times 10^{-18} \right\}^{\frac{1}{2}} \\ &= [2.364 \times 10^{-18}]^{\frac{1}{2}}, \end{aligned}$$

$$\hat{\sigma}_{y,TOTAL}(2s) = 1.54 \times 10^{-9}.$$

This value can be compared to the value obtained for the Allan deviation:

$$\begin{aligned} \sigma_y(2s) &= \left\{ \frac{1}{2} [\bar{y}_3 - \bar{y}_1]^2 \right\}^{\frac{1}{2}} \\ &= \left\{ \frac{1}{2} \left[\frac{x_5 - 2x_3 + x_1}{2} \right]^2 \right\}^{\frac{1}{2}} \\ &= \left\{ \frac{1}{8} [3.29 - 2(2.2) + 1.08]^2 \right\}^{\frac{1}{2}}, \end{aligned}$$

$$\sigma_y(2s) = 1.06 \times 10^{-11}.$$

$\sigma_y(2s)$ is seriously negatively biased by two orders of magnitude compared to $\hat{\sigma}_{y,TOTAL}(2s)$. (In the computation of the Allan deviation the OFD was not subtracted from the x_k s since the Allan deviation is insensitive to linear trends in the time domain).

A slight negative bias in $\hat{\sigma}_{y,TOTAL}(\tau)$ has been found for flicker FM noise and random walk FM noise. The bias for random walk FM noise can be removed by adjusting Eq. B5 [Howe, 1997].

B4. Other Variances. Several other variances have been introduced by workers in this field. In particular, before the introduction of the two-sample variance, it was standard practice to use the sample variance s^2 , defined as

$$s^2 = \int_0^{f_h} S_y(f) \left(\frac{\sin \pi f \tau}{\pi f \tau} \right)^2 df. \quad (\text{Eq. B9})$$

In practice it may be obtained from a set of measurements of the frequency of the oscillator as

$$s^2 = \frac{1}{N} \sum_{i=1}^N (y_i - \bar{y})^2. \quad (\text{Eq. B10})$$

The sample variance diverges for some types of noise and, therefore, is not generally useful.

Other variances based on the structure function approach can also be defined [Lindsey and Chi, 1976]. For example, there are the Hadamard variance, the three-sample variance and the high-pass variance [Rutman 1978]. They are occasionally used in research and scientific works for specific purposes, such as differentiating between different types of noise and for dealing with systematics and sidebands in the spectrum.

Appendix C

Confidence Intervals for $\sigma_y(\tau)$ - Examples

As discussed in section 4, the Allan variance has a chi-squared distribution function. Therefore, one way to obtain the confidence interval for a certain number of samples is to use

$$\chi^2 = (\text{df}) \frac{\hat{\sigma}_y^2}{\sigma_y^2}, \quad (\text{Eq. C1})$$

where df is the number of degrees of freedom (given in Table C1) [Howe, 1981]. This table is valid only for overlapped estimates of the Allan variance.

To compute the confidence interval for $\hat{\sigma}_y(\tau = 1 \text{ s}) = 10^{-12}$, for flicker frequency noise, $N = 101$, and $\tau_o = 0.5 \text{ s}$ ($m = 2$), we first find the number of degrees of freedom using Table C1:

$$\text{df} = \frac{5N^2}{4m(N+3m)} \quad (\text{Eq. C2})$$

$$= \frac{5(101^2)}{4(2)(101+3(2))} = 59.6.$$

For a 1σ (68%) confidence interval, the χ^2 values needed are $\chi^2(0.16)$ and $\chi^2(1-0.16)$. These values can be obtained from numerical tables of the chi-squared distribution function or from various computer programs. For this example $\chi^2(0.16) = 48.25$, and $\chi^2(0.84) = 69.73$.

Therefore

$$\chi^2(0.16) < \frac{59.6 \hat{\sigma}_y^2}{\sigma_y^2} < \chi^2(0.84), \quad (\text{Eq. C3})$$

or

$$\frac{59.6 \hat{\sigma}_y^2}{\chi^2(0.84)} < \sigma_y^2 < \frac{59.6 \hat{\sigma}_y^2}{\chi^2(0.16)}, \quad (\text{Eq. C4})$$

$$0.85 \hat{\sigma}_y^2 < \sigma_y^2 < 1.24 \hat{\sigma}_y^2. \quad (\text{Eq. C5})$$

Other methods have been developed for computing the confidence interval of Mod $\sigma_y(\tau)$ [Walter, 1994; Greenhall, 1995]. Table C2 shows a comparison of confidence intervals for $\sigma_y(\tau)$ (no overlap and full overlap) and Mod $\sigma_y(\tau)$ for white PM, flicker PM, and white FM noise processes [Walls, 1995]. As shown in Table C2, *overlapped* estimates improve the confidence intervals for specific values of M and m . Although $\sigma_y(\tau)$ usually provides a better estimate (confidence interval) than Mod $\sigma_y(\tau)$ for fractional fluctuations (see Table C2 for white PM and flicker PM), the confidence interval for absolute fluctuations is approximately similar. The reason is that the value of Mod $\sigma_y(\tau)$ is typically much smaller than $\sigma_y(\tau)$ for white and flicker PM noise processes; therefore Mod $\sigma_y(\tau)$ provides a better estimate of stability.

Table C1. Empirical equations for the number of degrees of freedom of the Allan Variance estimate. From [Howe, 1981].

Noise Type	Degrees of Freedom
white phase	$\frac{(N+1)(N-2m)}{2(N-m)}$
flicker phase	$\exp \left[\ln \left(\frac{N-1}{2m} \right) \ln \left(\frac{(2m+1)(N-1)}{4} \right) \right]^{\frac{1}{2}}$
white frequency	$\left[\frac{3(N-1)}{2m} - \frac{2(N-2)}{N} \right] \frac{4m^2}{4m^2+5}$
flicker frequency	$\frac{2(N-2)^2}{2.3N-4.9}$ for $m=1$ $\frac{5N^2}{4m(N+3m)}$ for $m \geq 2$
random-walk frequency	$\frac{N-2}{m} \frac{(N-1)^2 - 3m(N-1) + 4m^2}{(N-3)^2}$

Table C2. Confidence intervals for $\sigma_y(\tau)$ (no overlap and full overlap) and Mod $\sigma_y(\tau)$. From [Walls, 1995].

	No Overlap ± for 68% $\sigma_y(\tau)$	Full Overlap - for 68% $\sigma_y(\tau)$	Full Overlap + for 68% $\sigma_y(\tau)$	Full Overlap - for 68% Mod $\sigma_y(\tau)$	Full Overlap + for 68% Mod $\sigma_y(\tau)$
n=1025	White PM	White PM	White PM	White PM	White PM
m=2	4.4%	2.9%	3.2%	3.1%	3.4%
m=8	8.7%	2.9%	3.2%	5.2%	6.1%
m=32	17.4%	3.0%	3.4%	9.7%	14%
m=128	34.9%	3.1%	3.6%	18%	41%
n=1025	Flicker PM	Flicker PM	Flicker PM	Flicker PM	Flicker PM
m=2	4.4%	2.9%	3.1%	3.0%	3.3%
m=8	8.7%	3.6%	4.0%	5.7%	6.8%
m=32	17.4%	5.2%	6.1%	11%	16%
m=128	34.9%	8.4%	11%	20%	50%
n=1025	White FM	White FM	White FM	White FM	White FM
m=2	3.8%	2.8%	3.0%	3.0%	3.2%
m=8	7.7%	4.8%	5.6%	5.8%	7.0%
m=32	15.3%	8.8%	12%	11%	16%
m=128	30.6%	16%	32%	20%	51%

Appendix D

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NOTE: Appendix D does not repeat the references given in Section 7 of the main text.

Appendix E

Definitions

The standard definitions given in the main body of the text are repeated here in narrative form. If ambiguities are created between the narrative definition given here and the mathematical equation given in the text, the text has priority.

Frequency Deviation $y(t)$: instantaneous, normalized frequency departure from a nominal frequency.

Phase Deviation $\phi(t)$: instantaneous phase departure from a nominal phase.

Time Deviation $x(t)$: instantaneous time departure from a nominal time.

Frequency Instability $S_y(f)$: one-sided spectral density of the frequency deviation.

Phase Instability $S_\phi(f)$: one-sided spectral density of the phase deviation.

Time Instability $S_x(f)$: one-sided spectral density of the time deviation.

Two-Sample Variance, $\sigma_y^2(\tau)$, also called the Allan variance: time average over the sum of the squares of the differences between successive readings of the frequency departure sampled over the sampling time τ , under the assumption that there is no dead time between the frequency departure samples.

Two-Sample Deviation, $\sigma_y(\tau)$, also called the square root of the Allan variance: this is the square root of the two-sample variance as defined above.

Time Interval Error TIE: the variation of the time difference between a real clock and an ideal uniform time scale following a time period t after perfect synchronization.

Confidence Limit I_α : the uncertainty associated with the estimate of the two-sample deviation from a finite data set with a finite number of measurements. Typically, I_α refers to a 1σ (68%) confidence level or error bar. However, other confidence intervals, in particular larger ones, are frequently used.